

Consistency : الاتساق

Def:

Let $X_1, X_2, \dots, X_n$ be a random sample from a population with p.d.f. $f(x; \theta)$, $\theta \in \Omega$. The estimator $\hat{\theta}_n$ is called a consistent estimator for the parameter θ if and only if :-

$$\lim_{n \rightarrow \infty} E[\hat{\theta}_n - \theta] = 0 \quad \forall \theta \in \Omega$$

Which implies two conditions :-

- 1- $\lim_{n \rightarrow \infty} E(\hat{\theta}_n) = \theta$
- 2- $\lim_{n \rightarrow \infty} \text{Var}(\hat{\theta}_n) = 0$

* Example:

Let $X_1, X_2, \dots, X_n$ be a r.s. from a normal distribution with (μ, σ^2) . Are $\bar{X}$ and $S^2 = \frac{1}{n} \sum (x_i - \bar{x})^2$ consistent estimators for μ and σ^2 respectively?

Solu:

To check consistency for $\bar{X}$, we do have the following :-

- 1) $\lim_{n \rightarrow \infty} E \bar{X} = \lim_{n \rightarrow \infty} \mu = \mu$
- 2) $\lim_{n \rightarrow \infty} \text{var}(\bar{X}) = \lim_{n \rightarrow \infty} \frac{\sigma^2}{n} = 0$

Hence $\bar{X}$ is a consistent estimator for μ .

Now, for the consistency of S^2 .

$$\begin{aligned}
 1) \lim_{n \rightarrow \infty} E S^2 &= \lim_{n \rightarrow \infty} E \left[\frac{1}{n} \sum (x_i - \bar{X})^2 \right] = \lim_{n \rightarrow \infty} \left(\frac{n-1}{n} \sigma^2 \right) \\
 &\quad \text{بقسمة البسط والمقام على } n \\
 &= \lim_{n \rightarrow \infty} \left(\frac{\frac{n}{n} - \frac{1}{n}}{\frac{n}{n}} \sigma^2 \right) = \lim_{n \rightarrow \infty} \left(\frac{1 - \frac{1}{n}}{1} \sigma^2 \right) = \lim_{n \rightarrow \infty} \left(1 - \frac{1}{n} \right) \sigma^2 \\
 &= \sigma^2
 \end{aligned}$$

$$\begin{aligned}
 2) \lim_{n \rightarrow \infty} \text{Var}(S^2) &= \lim_{n \rightarrow \infty} \left[\frac{2(n-1)}{n^2} \right] \sigma^4 \\
 &\quad \text{بقسمة البسط والمقام على } n^2 \\
 &= \lim_{n \rightarrow \infty} \left[\frac{2}{n} - \frac{2}{n^2} \right] \sigma^4 \\
 &= 0
 \end{aligned}$$

Hence $S^2 = \frac{1}{n} \sum (x_i - \bar{X})^2$ is a consistent estimator for σ^2 .

* Example:

- 1- Is $\frac{x+1}{n+2}$ a consistent estimator for parameter of binomial distribution?
- 2- Is $\bar{X}$ is a consistent estimator for parameter of poisson distribution?
- 3- Is $(\bar{X} + 1)$ is a consistent estimator for mean of $N(\theta, \sigma^2)$?

Solu (1)

$$E\left[\frac{X+1}{n+2}\right] = \frac{EX+1}{n+2} = \frac{n\theta+1}{n+2}$$

$$1 - \lim_{n \rightarrow \infty} E\left[\frac{X+1}{n+2}\right] = \lim_{n \rightarrow \infty} \left[\frac{n\theta+1}{n+2}\right] = \lim_{n \rightarrow \infty} \frac{\theta + \frac{1}{n}}{1 + \frac{2}{n}} = \theta$$

$\therefore \frac{X+1}{n+2}$ is unbiased estimator for θ in limit

$$2 - \lim_{n \rightarrow \infty} \text{Var}\left[\frac{X+1}{n+2}\right] = \lim_{n \rightarrow \infty} \frac{\text{Var}(X)+0}{(n+2)^2} = \lim_{n \rightarrow \infty} \frac{n\theta(1-\theta)}{n^2+4n+4}$$

$$\text{بقسمة البسط والمقام على } n^2 = \lim_{n \rightarrow \infty} \frac{\frac{\theta(1-\theta)}{n}}{1 + \frac{4}{n} + \frac{4}{n^2}} = 0$$

$\therefore$ from (1) and (2) $\frac{X+1}{n+2}$ is consistent estimator for the parameter of binomial distribution.

Solu (2)

$$① \lim_{n \rightarrow \infty} E\bar{X} = \lim_{n \rightarrow \infty} E\left[\frac{1}{n} \sum_{i=1}^n x_i\right] = \lim_{n \rightarrow \infty} \frac{1}{n} \sum_{i=1}^n E x_i = \lim_{n \rightarrow \infty} \frac{n\lambda}{n} = \lambda$$

$$② \lim_{n \rightarrow \infty} \text{Var}(\bar{X}) = \lim_{n \rightarrow \infty} \text{Var}\left[\frac{1}{n} \sum x_i\right] = \lim_{n \rightarrow \infty} \left[\frac{1}{n^2} \sum \text{Var}(x_i)\right]$$

$$= \lim_{n \rightarrow \infty} \left[\frac{1}{n^2} n\lambda\right] = \lim_{n \rightarrow \infty} \frac{\lambda}{n} = 0$$

$\therefore \bar{X}$ is a consistent estimator for parameter of poisson distribution.

Solu (3)

$$① \lim_{n \rightarrow \infty} E(\bar{X}+1) = \lim_{n \rightarrow \infty} E\left[\frac{\sum x_i}{n} + 1\right] = \lim_{n \rightarrow \infty} \left[\frac{\sum E x_i}{n} + 1\right]$$

$$= \lim_{n \rightarrow \infty} \left[\frac{n\theta}{n} + 1\right] = \lim_{n \rightarrow \infty} [\theta + 1] = \theta + 1$$

$\therefore (\bar{X}+1)$ is non-consistent estimator for the mean of $N(\theta, \sigma^2)$

* Example:

Is $\bar{X}$ a consistent estimator for θ of :-

- 1) If $x_1, \dots, x_n \sim \text{Ber}(\theta)$
- 2) If $x_1, \dots, x_n \sim \text{bin}(n, \theta)$

Solu:

$$\textcircled{1} f(x; \theta) = \theta^x (1-\theta)^{1-x}$$

$$1) \lim_{n \rightarrow \infty} E(\bar{X}) = \lim_{n \rightarrow \infty} E\left[\frac{\sum x_i}{n}\right] = \lim_{n \rightarrow \infty} \frac{\sum_{i=1}^n E x_i}{n} = \lim_{n \rightarrow \infty} \frac{\sum_{i=1}^n \theta}{n}$$

$$= \lim_{n \rightarrow \infty} \frac{n\theta}{n} = \lim_{n \rightarrow \infty} \theta = \theta$$

$$2) \lim_{n \rightarrow \infty} \text{Var}(\bar{X}) = \lim_{n \rightarrow \infty} \text{Var}\left(\frac{\sum x_i}{n}\right) = \lim_{n \rightarrow \infty} \frac{1}{n^2} \sum_{i=1}^n \text{Var}(x_i)$$

$$= \lim_{n \rightarrow \infty} \frac{1}{n^2} \sum_{i=1}^n [\theta(1-\theta)] = \lim_{n \rightarrow \infty} \frac{n\theta(1-\theta)}{n^2}$$

$$= \lim_{n \rightarrow \infty} \frac{\theta(1-\theta)}{n} = 0$$

$\therefore \bar{X}$ is a consistent estimator for θ (the parameter of Bernoulli distribution).

$$\textcircled{2} f(x; n, \theta) = C_x^n \theta^x (1-\theta)^{n-x} \quad x = 0, 1, 2, \dots, n$$

$$1- \lim_{n \rightarrow \infty} E\bar{X} = \lim_{n \rightarrow \infty} E\left(\frac{X}{n}\right) ; X = \sum_{i=1}^n x_i = \text{no. of successes}$$

$\bar{X} = \text{proportion of successes}$

$$\lim_{n \rightarrow \infty} E\bar{X} = \lim_{n \rightarrow \infty} \frac{EX}{n} = \lim_{n \rightarrow \infty} \frac{n\theta}{n} = \lim_{n \rightarrow \infty} \theta = \theta$$

$$2- \lim_{n \rightarrow \infty} \text{Var}(\bar{X}) = \lim_{n \rightarrow \infty} \text{Var}\left(\frac{X}{n}\right) = \lim_{n \rightarrow \infty} \frac{\text{Var}(X)}{n^2}$$

$$= \lim_{n \rightarrow \infty} \frac{n\theta(1-\theta)}{n^2} = \lim_{n \rightarrow \infty} \frac{\theta(1-\theta)}{n} = 0$$

$\therefore \bar{X}$ is a consistent estimator for θ (the parameter of binomial distribution).