

\* Example:

Let  $X_1, X_2, \dots, X_n$  be a r.s. of size  $n$  from Bernoulli distribution with parameter  $\theta$ , and let  $\bar{X}$  be an estimator for  $\theta$ . Apply Cramer-Rao inequality on the  $\text{Var}(\bar{X})$ , and find the lower bound (L.B.) for  $\text{Var}(\bar{X})$

Solu:

$$f(x; \theta) = \theta^x (1-\theta)^{1-x} \quad x=0, 1$$

$$\ln f(x; \theta) = x \ln \theta + (1-x) \ln (1-\theta)$$

$$\frac{\partial \ln f(x; \theta)}{\partial \theta} = \frac{x}{\theta} - \frac{1-x}{1-\theta}$$

$$\frac{\partial^2 \ln f(x; \theta)}{\partial \theta^2} = -\frac{x}{\theta^2} - \frac{1-x}{(1-\theta)^2}$$

$$\begin{aligned} I_x^{(\theta)} &= -E\left[\frac{\partial^2 \ln f(x; \theta)}{\partial \theta^2}\right] = \frac{E x}{\theta^2} + \frac{E(1-x)}{(1-\theta)^2} \\ &= \frac{\theta}{\theta^2} + \frac{1-\theta}{(1-\theta)^2} = \frac{1}{\theta} + \frac{1}{1-\theta} = \frac{1-\theta+\theta}{\theta(1-\theta)} = \frac{1}{\theta(1-\theta)} \end{aligned}$$

$$\Rightarrow \text{Fisher Information} = n I_x^{(\theta)} = \frac{n}{\theta(1-\theta)}$$

$$E(\bar{X}) = E\left[\frac{\sum x_i}{n}\right] = \frac{1}{n} \sum_{i=1}^n E x_i = \frac{1}{n} n \theta = \theta$$

$\therefore \bar{X}$  is unbiased estimator for  $\theta$ .

We have Cramer-Rao inequality

$$\text{Var}(T) \geq \frac{[g'(\theta)]^2}{n I_x^{(\theta)}} \Rightarrow \text{Var}(\bar{X}) \geq \frac{1}{n I_x^{(\theta)}}$$

$$\Rightarrow \text{Var}(\bar{X}) \geq \frac{\theta(1-\theta)}{n}$$

$\therefore$  The lower bound for  $\text{Var}(\bar{X})$  is

$$\text{Var}(\bar{X}) = \frac{\theta(1-\theta)}{n}$$

Note:

If  $\hat{\theta}$  is the unbiased estimator for  $\theta$  and

$$\text{Var}(\hat{\theta}) = \frac{[g'(\theta)]^2}{n I_x(\theta)}$$

Then we say that  $\hat{\theta}$  is a minimum variance unbiased estimator for  $(\theta)$  (m.v.u.e)

$\therefore$  From the above example

$$\therefore E(\bar{X}) = \theta \text{ and } \text{Var}(\bar{X}) = \frac{1}{n I_x(\theta)}$$

$\therefore \bar{X}$  is a m.v.u.e for  $\theta$ .

\* Example:

Let  $X_1, X_2, \dots, X_n$  be a r.s. of size  $n$  from  $N(\theta, \sigma^2)$  and let  $\bar{X}$  be an estimator for  $\theta$ . Find C.R.L.B for  $\text{Var}(\bar{X})$ .

Solu.

$$f(x; \theta, \sigma^2) = \frac{1}{\sqrt{2\pi} \sigma} e^{-\frac{1}{2\sigma^2}(x-\theta)^2}$$

$$\ln f(x; \theta, \sigma^2) = \ln \left[ \frac{1}{\sqrt{2\pi} \sigma} \right] - \frac{1}{2\sigma^2}(x-\theta)^2$$

$$\frac{\partial \ln f(x; \theta, \sigma^2)}{\partial \theta} = \frac{-2}{2\sigma^2} (x - \theta)(-1) = \frac{x - \theta}{\sigma^2} = \frac{x}{\sigma^2} - \frac{\theta}{\sigma^2}$$

$$\frac{\partial^2 \ln f(x; \theta)}{\partial \theta^2} = \frac{-1}{\sigma^2}$$

$$I_x(\theta) = -E\left[\frac{-1}{\sigma^2}\right] = \frac{1}{\sigma^2}$$

Fisher information in  $n$  random observations

$$= n I_x(\theta) = \frac{n}{\sigma^2}$$

From C.R. inequality we have:

$$\text{Var}(T) \geq \frac{[g'(\theta)]^2}{n I_x(\theta)}$$

$$T = \bar{X}, g(\theta) = \theta \Rightarrow g'(\theta) = 1$$

$$\text{C.R.L.B for } \text{Var}(\bar{X}) = \frac{1}{n I_x(\theta)} = \frac{1}{\frac{n}{\sigma^2}} = \frac{\sigma^2}{n}$$

\* from the above example is  $\bar{X}$  a m.v.u.e for  $\theta$ ?

Solu.  $\because x \sim N(\theta, \sigma^2)$

$$\Rightarrow \bar{X} \sim N\left(\theta, \frac{\sigma^2}{n}\right) \Rightarrow \text{Var}(\bar{X}) = \frac{\sigma^2}{n} = \text{C.R.L.B.}$$

$$\text{and } E(\bar{X}) = E\left[\frac{\sum x_i}{n}\right] = \frac{\sum E(x_i)}{n} = \frac{n\theta}{n} = \theta \text{ (unbiased)}$$

$\therefore \bar{X}$  is a m.v.u.e for  $\theta$ .

\* Example:

Let  $x_1, x_2, \dots, x_n$  be a r.s. of size  $n$  from poisson distribution with parameter  $\theta$  and let  $\bar{X}$  be an estimator for  $\theta$ . Is  $\bar{X}$  a m.v.u.e for  $\theta$ ?

Solu:

$$f(x; \theta) = \frac{e^{-\theta} \theta^x}{x!}$$

$$\ln f(x; \theta) = -\theta + x \ln \theta - \ln x!$$

$$\frac{\partial \ln f(x; \theta)}{\partial \theta} = -1 + \frac{x}{\theta}$$

$$\frac{\partial^2 \ln f(x; \theta)}{\partial \theta^2} = -\frac{x}{\theta^2}$$

$$\Rightarrow I_x^{(\theta)} = -E\left[\frac{-x}{\theta^2}\right] = \frac{Ex}{\theta^2} = \frac{\theta}{\theta^2} = \frac{1}{\theta}$$

$\Rightarrow$  Fisher information in  $n$  random observations

$$n I_x^{(\theta)} = \frac{n}{\theta}$$

$$T = \bar{X}, g(\theta) = \theta \Rightarrow g'(\theta) = 1$$

$$\Rightarrow \text{C.R.L.B for variance of } \bar{X} = \frac{1}{n I_x^{(\theta)}} = \frac{1}{\frac{n}{\theta}} = \frac{\theta}{n}$$

$$\text{Var}(\bar{X}) = \text{Var}\left(\frac{\sum x_i}{n}\right) = \frac{n \text{Var}(x_i)}{n^2} = \frac{n\theta}{n^2} = \frac{\theta}{n} = \text{C.R.L.B for Var}(\bar{X})$$

$$\therefore E(\bar{X}) = E\left[\frac{\sum x_i}{n}\right] = \frac{n E x_i}{n} = \frac{n\theta}{n} = \theta \text{ (unbiased)}$$

$\therefore \bar{X}$  is a m.v.u.e for  $\theta$ .