Mathematical Statistics II - Spring 2024 Lecture 03

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Poisson distribution: A random variable X is said to have a Poisson distribution, denoted by $X \sim Poisson(p)$, if its PMF has the following form:

$$f(x) = \begin{cases} \frac{\lambda^x e^{-\lambda}}{x!} & \text{for } x = 0, 1, 2, \dots \\ 0 & \text{otherwise} \end{cases}$$

where $\lambda > 0$ is the expected rate of occurrence of a certain event. This probability distribution models the behavior of a random variable X that represents the number of times a certain event occurs over a specified period. For instance, the number of monthly car accidents in a certain city, the number of daily COVID cases in Iraq, and the number of defective items per container produced by a specific machine.

Mean, variance, MGF, and CDF: by definition of the mean, variance, and MGF are defined as follows:

$$E[X] = \sum_{x=0}^{\infty} x \frac{\lambda^x e^{-\lambda}}{x!}$$
$$= \sum_{x=1}^{\infty} x \frac{\lambda^x e^{-\lambda}}{x!}$$
$$= \lambda e^{-\lambda} \sum_{x=1}^{\infty} \frac{\lambda^{x-1}}{(x-1)!}$$

let k = x - 1, then k = 0, 1, 2, ...

$$E[X] = \lambda e^{-\lambda} \sum_{k=0}^{\infty} \frac{\lambda^k}{k!}$$
$$= \lambda e^{-\lambda} e^{\lambda} = \lambda$$

to find $E[X^2]$, we obtain E[X(X-1)] first. Then calculate $E[X^2] = E[X(X-1)] + E[X]$

$$\begin{split} E[X(X-1)] &= \sum_{x=0}^{\infty} x(x-1) \frac{\lambda^x e^{-\lambda}}{x!} \\ &= \sum_{x=2}^{\infty} x(x-1) \frac{\lambda^x e^{-\lambda}}{x!}, \text{ when } x=0 \text{ the first two terms are zeros} \\ &= \lambda^2 e^{-\lambda} \sum_{x=2}^{\infty} \frac{\lambda^x}{(x-1)!} \end{split}$$

let k = x - 2, then k = 0, 1, 2, ...

$$E[X(X-1)] = \lambda^2 e^{-\lambda} \sum_{k=0}^{\infty} \frac{\lambda^k}{k!}$$
$$= \lambda^2 e^{-\lambda} e^{\lambda} = \lambda^2$$

$$E[X^{2}] = E[X(X - 1)] + E[X]$$
$$= \lambda^{2} + \lambda$$

$$Var(X) = E[X^{2}] - E[X]^{2}$$
$$= \lambda^{2} + \lambda - \lambda^{2} = \lambda$$

notice that the mean and variance of Poisson random variable are equal.

$$M_X(t) = \sum_{x=0}^{\infty} e^{tx} \frac{\lambda^x e^{-\lambda}}{x!}$$
$$= e^{-\lambda} \sum_{x=0}^{\infty} \frac{(\lambda e^t)^x}{x!}$$
$$= e^{-\lambda} e^{\lambda e^t} = e^{\lambda(e^t - 1)}$$

Example 3.1: Let X be a random variable that follows a Poisson distribution with rate λ . Find the value of λ , such that P(X = 0) = P(X = 1). Then find E[X], Var(X), and $M_X(t)$.

Solution: To find the value of λ , we have:

$$P(X = 0) = P(X = 1)$$

$$\Rightarrow \frac{\lambda^0 e^{-\lambda}}{0!} = \frac{\lambda^1 e^{-\lambda}}{1!}$$

$$\Rightarrow e^{-\lambda} = \lambda e^{-\lambda}$$

$$\Rightarrow \lambda = \frac{e^{-\lambda}}{e^{-\lambda}} = 1$$

hence the PMF of X is

$$f(x) = \begin{cases} \frac{e^{-1}}{x!} & \text{for } x = 0, 1, 2, \dots \\ 0 & \text{otherwise} \end{cases}$$

therefore E[X] = Var(X) = 1, and $M_X(t) = e^{e^t - 1}$.

Example 3.2: Let $X_1 \sim Poisson(4), X_2 \sim Poisson(6), \text{ and } X_3 \sim$

Poisson(8) are three independent random variables. Do the following:

1- Find the mean and variance of $Y_1 = X_1 + X_2$, $Y_2 = X_1 + X_3$, and $Y_3 = X_1 + X_2 + X_3$.

- 2- The MGF of $Y_3 = X_1 + X_2 + X_3$.
- 3- Find $P(Y_1 \ge 2)$.

Solution: 1- Find the mean and variance

$$E[Y_1] = E[X_1 + X_2]$$
$$= E[X_1] + E[X_2] = 4 + 6 = 10$$

$$E[Y_2] = E[X_1 + X_3]$$
$$= E[X_1] + E[X_3] = 4 + 8 = 12$$

$$E[Y_3] = E[X_1 + X_2 + X_3]$$

$$= E[X_1] + E[X_2] + E[X_3] = 4 + 6 + 8 = 18$$

$$Var(Y_1) = Var(X_1 + X_2)$$

= $Var(X_1) + Var(X_2) = 4 + 6 = 10$

$$Var(Y_2) = Var(X_1 + X_2)$$

= $Var(X_1) + Var(X_3) = 4 + 8 = 12$

$$Var(Y_3) = Var(X_1 + X_2 + X_3)$$

= $Var(X_1) + Var(X_2) + Var(X_3) = 4 + 6 + 8 = 18$

we will see later this semester that $Y_1 \sim Poisson(\lambda_1 + \lambda_2)$, $Y_2 \sim Poisson(\lambda_1 + \lambda_3)$, and $Y_3 \sim Poisson(\lambda_1 + \lambda_2 + \lambda_3)$.

2- Find MGF of
$$Y_3 = X_1 + X_2 + X_3$$

$$M_{Y_3}(t) = E \left[e^{tY_3} \right] = E \left[e^{t(X_1 + X_2 + X_3)} \right]$$

$$= E \left[e^{tX_1} e^{tX_2} e^{tX_3} \right]$$

$$= E \left[e^{tX_1} \right] E \left[e^{tX_2} \right] E \left[e^{tX_3} \right]$$

$$= M_{X_1}(t) M_{X_2}(t) M_{X_3}(t)$$

$$= e^{\lambda_1(e^t - 1)} e^{\lambda_2(e^t - 1)} e^{\lambda_3(e^t - 1)}$$

$$= e^{4(e^t - 1)} e^{6(e^t - 1)} e^{8(e^t - 1)}$$

$$= e^{(4 + 6 + 8)(e^t - 1)} = e^{18(e^t - 1)}$$

$$P(Y_1 \ge 2) = 1 - P(Y_1 \le 1)$$

$$= 1 - P(X_1 + X_2 \le 1)$$

$$= 1 - (P(X_1 = 0, X_1 = 0) + P(X_1 = 0, X_2 = 1) + P(X_1 = 1, X_2 = 0))$$

$$= 1 - \left(\frac{4^0 e^{-4}}{0!} \times \frac{6^0 e^{-6}}{0!} + \frac{4^0 e^{-4}}{0!} \times \frac{6^1 e^{-6}}{1!} + \frac{4^1 e^{-4}}{1!} \times \frac{6^0 e^{-6}}{0!}\right)$$

$$= 1 - \left(e^{-10} + 6e^{-10} + 4e^{-10}\right)$$

$$= 1 - 11e^{-10} = 0.9995$$

Now assume that $Y_1 \sim Poisson(10)$, find $P(Y_1 \ge 2)$. (**Homework**)

Poisson distribution as an approximation to binomial distribution

Let X follows a binomial distribution with the total number of trials n and the probability of success p. If n is very large (say $n \ge 100$), and p is small, such that $np \le 10$, then the distribution of X can be approximated using the Poisson distribution with rate $\lambda = np$. (**Proof is Homework**) **Example 3.3**: Suppose that 0.003 of bolts made by a machine are defective, the defectives occurring at random during production. If the bolts are packaged in boxes of 100, what is the Poisson approximation that a given box will contain x defectives?

Solution: Let X represents the number of defective bolts in each box. Then X has a binomial distribution with n = 100 and p = 0.003. Now,

since

$$np = 100(0.003) = 0.3 \le 10$$

we can use Poisson distribution with rate $\lambda = np = 0.3$ to approximate the distribution of X. That is:

$$f(x) \approx \frac{\lambda^x e^{-\lambda}}{x!}$$
$$= \frac{0.3^x e^{-0.3}}{x!}$$

for x = 0, 1, 2, ...

Example 3.4 (Homework): Given that 0.0004 of vehicles break down when driving through a certain tunnel find the probability of (a) no car breaks down (b) at least two cars break down in an hour when 2,000 vehicles enter the tunnel.

Example 3.5: Let $X \sim Poisson(\lambda)$, and assume that the conditional distribution of Y given X = x is Binomial(x, p). Show that $Y \sim Poisson(p\lambda)$.

Solution: We have

$$f_X(x) = \begin{cases} \frac{\lambda^x e^{-\lambda}}{x!} & x = 0, 1, 2, \dots \\ 0 & \text{otherwise} \end{cases}$$

$$f_{Y|X}(y|x) = \begin{cases} \binom{x}{y} p^y (1-p)^{x-y} & y = 0, 1, ..., x; \text{ and } 0 \le p \le 1\\ 0 & \text{otherwise} \end{cases}$$

to find the marginal distribution of Y, we first need to find the joint distribution of X and Y. By definition, the conditional PMF can be written as

$$f_{Y|X}(y|x) = \frac{f_{X,Y}(x,y)}{f_X(x)}$$

$$\implies f_{X,Y}(x,y) = f_{Y|X}(y|x)f_X(x)$$

$$= \binom{x}{y}p^y(1-p)^{x-y}\frac{\lambda^x e^{-\lambda}}{x!}$$

$$= \frac{x!}{y!(x-y)!}p^y(1-p)^{x-y}\frac{\lambda^x e^{-\lambda}}{x!}$$

$$= \frac{p^y e^{-\lambda}}{y!} \times \frac{\lambda^x (1-p)^{x-y}}{(x-y)!}$$

for $y = 0, 1, ..., x; x \ge y; 0 \le p \le 1$; and $\lambda > 0$.

Now to find the marginal PMF of Y we need to sum over x. That is:

$$f_Y(y) = \sum_{x=y}^{\infty} \frac{p^y e^{-\lambda}}{y!} \times \frac{\lambda^x (1-p)^{x-y}}{(x-y)!}$$
$$= \frac{p^y e^{-\lambda}}{y!} \left(\sum_{x=y}^{\infty} \frac{\lambda^x (1-p)^{x-y}}{(x-y)!} \right)$$

let k = x - y, then x = k + y, and k = 0, 1, 2, ...

$$f_Y(y) = \frac{p^y e^{-\lambda}}{y!} \left(\sum_{k=0}^{\infty} \frac{\lambda^{k+y} (1-p)^k}{k!} \right)$$

$$= \frac{p^y e^{-\lambda} \lambda^y}{y!} \left(\sum_{k=0}^{\infty} \frac{((1-p)\lambda)^k}{k!} \right)$$

$$= \frac{(p\lambda)^y e^{-\lambda} e^{(1-p)\lambda}}{y!}$$

$$= \frac{(p\lambda)^y e^{-\lambda} e^{\lambda} e^{-p\lambda}}{y!} = \frac{(p\lambda)^y e^{-p\lambda}}{y!}$$

the later is the PMF of Poisson distribution with rate parameter $p\lambda$. Therefore, $Y \sim Poisson(p\lambda)$.