

Lecture (3)

- Some special mathematical expectation
- Examples

Remark:-

From Ex(3), we can conclude that in general, the expected value of a product of functions is not equal to the product of the expected value of the same functions.

Ex(4):- واجب

Consider the following probability distribution table:-

x	0	1	2
$f(x)$	$5/12$	$3/12$	$4/12$

Find the expected value of x

Sol:-

$$E(x) = \sum_x f(x) \cdot x \quad \text{or} \quad \sum_x x p(x)$$

$$= \sum_{x=0}^2 x p(x) = 0 p(X=0) + 1 p(X=1) + 2 p(X=2)$$

$$= 0 + 1\left(\frac{3}{12}\right) + 2\left(\frac{4}{12}\right) = \frac{11}{12}$$

Ex(5):- واجب

Given a r.v. x with a p.d.f $f(x) = \begin{cases} e^{-x} & x > 0 \\ 0 & \text{o.w.} \end{cases}$

Find $E(x)$

Sol:-

$$E(x) = \int_x x f(x) dx = \int_0^{\infty} x e^{-x} dx$$

$$\text{let } u = x \Rightarrow du = dx$$

$$\text{and } dv = e^{-x} dx \Rightarrow v = \int e^{-x} dx = -e^{-x}$$

$$\int u dv = uv - \int v du$$

$$\Rightarrow \int_0^{\infty} x e^{-x} dx = -x e^{-x} \Big|_0^{\infty} - \int_0^{\infty} -e^{-x} dx$$

$$= -x e^{-x} \Big|_0^{\infty} - e^{-x} \Big|_0^{\infty}$$

$$= -\infty e^{-\infty} + 0 e^{-0} - e^{-\infty} + e^{-0} = 0 + 0 - 0 + 1 = 1$$

Ex(6)

Given a r.v. with a p.d.f $f(x) = \begin{cases} 2(1-x) & 0 < x < 1 \\ 0 & \text{o.w.} \end{cases}$

Show that $E(X^r) = \frac{2}{(r+1)(r+2)}$ and use it to

evaluate $E[(2X+1)^2]$

Sol:-

$$E(X^r) = \int_x x^r f(x) dx$$

$$= \int_0^1 x^r \cdot 2(1-x) dx = 2 \int_0^1 (x^r - x^{r+1}) dx$$

$$= 2 \left[\int_0^1 x^r dx - \int_0^1 x^{r+1} dx \right] = 2 \left[\frac{x^{r+1}}{r+1} - \frac{x^{r+2}}{r+2} \right] \Big|_0^1$$

$$= 2 \left[\frac{1}{r+1} - \frac{1}{r+2} \right] = 2 \left[\frac{r+2 - (r+1)}{(r+1)(r+2)} \right] = \frac{2}{(r+1)(r+2)}$$

Now:-

$$\begin{aligned} E[(2X+1)^2] &= E[(4X^2 + 4X + 1)] = 4E(X^2) + 4E(X) + 1 \\ &= 4 \cdot \frac{2}{(2+1)(2+2)} + 4 \cdot \frac{2}{(1+1)(1+2)} + 1 = \frac{8}{12} + \frac{8}{6} + 1 \\ &= \frac{36}{12} = 3 \end{aligned}$$

H.W:-

Given $f(x) = \begin{cases} 2e^{-x} & x > 0 \\ 0 & \text{o.w.} \end{cases}$

Find $E(X)$, $E(X^2)$, $E(2X+3X^2)$.

Ex(7):-

Consider the following p.m.f:-

$$f(x) = \begin{cases} \left(\frac{1}{3}\right)^x \left(\frac{2}{3}\right)^{1-x} & x = 0, 1 \\ 0 & \text{o.w.} \end{cases}$$

Find $E(X)$, $E(2X+3)$, $E(X^2)$, $E(X^3)$,
 $E(1+2X+3X^2+5X^3)$

Sol:-

$$\begin{aligned} E(X) &= \sum_x x f(x) = \sum_{x=0}^1 x \left(\frac{1}{3}\right)^x \left(\frac{2}{3}\right)^{1-x} \\ &= 0 \left(\frac{1}{3}\right)^0 \left(\frac{2}{3}\right) + 1 \left(\frac{1}{3}\right) \left(\frac{2}{3}\right)^0 = 0 + \frac{1}{3} = \frac{1}{3} \end{aligned}$$

$$E(2X+3) = 2E(X) + 3 = 2\left(\frac{1}{3}\right) + 3 = \frac{11}{3}$$

$$\begin{aligned} E(X^3) &= \sum_{x=0}^1 x^3 \left(\frac{1}{3}\right)^x \left(\frac{2}{3}\right)^{1-x} \\ &= (0)^3 \left(\frac{1}{3}\right)^0 \left(\frac{2}{3}\right) + (1)^3 \left(\frac{1}{3}\right) \left(\frac{2}{3}\right)^0 = \frac{1}{3} \end{aligned}$$

Some special mathematical expectation :-

1) The mean :-

If we let $U(x) = X$ then $E[U(x)] = E(X)$. This expectation, when it exists, is called the arithmetic mean or simply, the mean of X .

The mean of X denoted by μ is defined to be;

$$\mu = E(X) = \begin{cases} \sum_{\text{all } x} x f_x(x) & \text{if } X \text{ is a discrete r.v.} \\ \int_{-\infty}^{\infty} x f_x(x) dx & \text{if } X \text{ is a continuous r.v.} \end{cases}$$

2) The moments of a random variable:-

The moments (or raw moments) of a r.v. or a distribution are the expectation of the powers of the r.v. which has the given distribution.

Three kinds of moments are:-

i. The non-central moments or the moments about the origin:-

If X is a r.v., the m th moment of X , denoted by μ'_m is defined as:-

$$\mu'_m = E(X^m)$$

ii. The central moments :-

If X is a r.v., the m th central moment of X , about μ_x , denoted by

$$\mu_m = E(X - \mu_x)^m$$

and the m -th non-central moment of X is

defined as $E(X-a)^m$.

We see that the moment about the origin is a special case of non-central moment when $a=0$ and ~~also~~ the central moment is also the special case of non central moment when $a = \mu_x$.

A final note on raw and central moments can be summarized in the following two points;

a) If one of the two moments $E(X^m)$ and $E(X-\mu_x)^m$ exists, then the other moments of some order exists as well as moments of order less than m . That is, if $E(X^m)$ exists then $E(X-\mu_x)^m$ exists for all $n \leq m$ and Vice versa.

b) If one of the two moments $E(X^m)$ or $E(X-\mu_x)^m$ of a particular order does not exist, then the other moment of the same order does not exist and neither any other moment of order greater than m . That is, if $E(X^m)$ does not exist, then $E(X-\mu_x)^m$ does not exist for all $n \geq m$.

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If $E(X^2)$ exist, then $E(X)$ exist, also $E(X-\mu_x)^2$ exist and also $E(X-\mu_x)$. But if $E(X)$ does not exist, then neither $E(X-\mu_x)$ nor $E(X^2)$, $E(X^3)$, --- or $E(X-\mu_x)^2$, $E(X-\mu_x)^3$ --- exist,

iii) The third kind of moments are called factorial moments,

The m th factorial moment of a r.v. X (often discrete) is defined as:

$E[X(X-1)(X-2) \dots (X-m+1)]$ where m is a positive integer, number

3) The variance:-

Another special mathematical expectation is obtained by taking $U(x) = (X - \mu_x)^2$. The expected value of this function is called the variance of the r.v. X , and denoted by σ_x^2 or $\text{Var}(X)$. That is,

$$\sigma_x^2 = E(X - \mu_x)^2$$

$\sigma_x = \sqrt{\sigma_x^2}$ is called the standard deviation of the r.v. X , or of the distribution of the r.v. X .

Some theorems about the variance of a r.v. X are:-

Theorem (1):-

$$\text{Var}(X) \geq 0$$

proof:-

By definition

$$\therefore \text{Var}(X) = E(X - \mu_x)^2 \geq 0$$

$$\therefore \text{Var}(X) \geq 0,$$

Theorem (2):-

$$\text{Var}(C) = 0, \text{ where } C \text{ is a constant}$$

proof:-

$\text{Var}(C) = E[C - E(C)]^2$, but since C is a constant, then $E(C) = C$ by the fact of expectation and therefore

$$\text{Var}(C) = E(C - C)^2 = 0$$

theorem (3):-

$$\text{Var}(CX) = C^2 \text{Var}(X)$$

proof:-

$$\begin{aligned} \text{Var}(CX) &= E[CX - E(CX)]^2 \\ &= E[CX - C\mu_X]^2 \\ &= E[C^2 (X - \mu_X)^2] = C^2 E(X - \mu_X)^2 = C^2 \text{Var}(X) \end{aligned}$$

theorem (4):-

$$\sigma_X^2 = EX^2 - (E(X))^2 = EX^2 - \mu_X^2$$

proof:-

$$\begin{aligned} \sigma_X^2 &= E(X - \mu_X)^2 = E(X^2 - 2X\mu_X + \mu_X^2) \\ &= EX^2 - 2\mu_X EX + \mu_X^2 \\ &= EX^2 - 2\mu_X^2 + \mu_X^2 = EX^2 - \mu_X^2 \end{aligned}$$

Ex(1):-

Let a r.v. X has the following p.d.f

$$f(x) = \begin{cases} \frac{4-x}{16} & -2 < x < 2 \\ 0 & \text{o.w.} \end{cases}$$

- 1- Find the third non-central moment.
- 2- Find the third central moment about 4.
- 3- Find the second central moment about the mean

sol:-

$$\begin{aligned} 1) \mu_3' &= E(X^3) = \int_{-2}^2 X^3 f(x) dx = \int_{-2}^2 X^3 \frac{(4-X)}{16} dx \\ &= \frac{1}{4} \int_{-2}^2 X^3 dx - \frac{1}{16} \int_{-2}^2 X^4 dx \end{aligned}$$

$$\begin{aligned} &= \frac{1}{16} \left(X^4 \Big|_{-2}^2 \right) - \frac{1}{16(5)} \left(X^5 \Big|_{-2}^2 \right) = \frac{1}{16} \left(16 - \frac{1}{16} \right) - \frac{1}{80} \\ &\quad \cdot (32 + 32) = 0.8 \end{aligned}$$

$$\begin{aligned}
 2- E(X-4)^3 &= \int_{-2}^2 (X-4)^3 \cdot \frac{(4-X)}{16} dx \\
 &= -\frac{1}{16} \int_{-2}^2 (X-4)^4 dx = -\frac{1}{16} \cdot \frac{(X-4)^5}{5} \Big|_{-2}^2 \\
 &= -\frac{1}{16} \cdot \frac{(-2)^5}{5} + \frac{1}{16} \cdot \frac{(-6)^5}{5} = 0.4 - 97.2 = -96.8
 \end{aligned}$$

$$3- \sigma_2^2 = E(X - \mu_X)^2 = \sigma_X^2$$

$$\begin{aligned}
 EX &= \int_{-2}^2 X \cdot \frac{(4-X)}{16} dx = \frac{1}{16} \int_{-2}^2 (4X - X^2) dx \\
 &= \frac{1}{16} \left[2X^2 \Big|_{-2}^2 - \frac{X^3}{3} \Big|_{-2}^2 \right] = \frac{1}{16} \left[8 - 8 - \left(\frac{8}{3} + \frac{8}{3} \right) \right] \\
 &= -\frac{1}{3}
 \end{aligned}$$

$$\therefore E(X - \mu)^2 = E\left(X + \frac{1}{3}\right)^2$$

$$\begin{aligned}
 &= \int_{-2}^2 \left(X + \frac{1}{3}\right)^2 \cdot \frac{(4-X)}{16} dx \\
 &= \int_{-2}^2 \left(X^2 + \frac{2}{3}X + \frac{1}{9}\right) \cdot \frac{(4-X)}{16} dx \\
 &= \frac{1}{16} \left[4 \int_{-2}^2 X^2 dx + \frac{8}{3} \int_{-2}^2 X dx + \frac{4}{9} \int_{-2}^2 dx - \int_{-2}^2 X^3 dx - \frac{2}{3} \int_{-2}^2 X^2 dx - \frac{1}{9} \int_{-2}^2 X dx \right] \\
 &= \frac{1}{16} \left[4 \left(\frac{X^3}{3} \Big|_{-2}^2 \right) + \frac{4}{3} \left(X^2 \Big|_{-2}^2 \right) + \frac{2}{9} \left(X \Big|_{-2}^2 \right) - \left(\frac{X^4}{4} \Big|_{-2}^2 \right) - \frac{2}{9} \left(X^3 \Big|_{-2}^2 \right) - \frac{1}{18} \left(X^2 \Big|_{-2}^2 \right) \right] = 18.67
 \end{aligned}$$