

Mathematical Statistics

Fourth class

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الاحصاء الرياضي

المرحلة الرابعة

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$F(x)$: distribution function

دالة توزيع

متقطع

$$(1) F(x) = P(X \leq x) = \sum_{u \leq x} p_u \quad \text{if } x \text{ is discrete}$$

متنوع

$$(2) F(x) = \int_{-\infty}^x f(s) ds \quad \text{if } x \text{ is continuous}$$

$$\text{And } F(0) = 0 \quad F(\infty) = 1$$

$f(x)$:- Probability density function (P.D.F)

satisfies two conditions:-

2 شروط

$$(1) f(x) \geq 0 \quad \left. \vphantom{f(x) \geq 0} \right\} x \text{ is discrete}$$

$$(2) \sum_{x \in R_x} f(x) = 1$$

and

متنوع

$$(1) f(x) \geq 0 \quad \left. \vphantom{f(x) \geq 0} \right\} x \text{ is continuous}$$

$$(2) \int_{x \in R_y} f(x) dx = 1$$

(Mathematical Expectation)

التوقع الرياضي

$E(X) = M = \text{توقع} = \text{معدل} = \frac{\text{مجموع}}{n}$: first moment

$$E(X) = \sum_{x \in R_Y} x f(x)$$

disc

متقطع

$$E(X) = \int_{x \in R_Y} x f(x) dx$$

cont

مستمر

$$E(X) = M'_X(0) \Leftarrow E(X) = \text{first moment}$$

$$M''_X(0) \Leftarrow E(X^2) = \text{second moment}$$

$$\vdots M_X^{(r)}(0) \Leftarrow E(X^r) = r\text{-th moment}$$

التوزيع المشترك (Joint Distribution)

والمتغيرات العشوائية

if X and Y are two r.v.s then the joint P.f.f is given by $f(x,y)$ which satisfies the

(1) Where X, Y discrete $\text{عندما } X, Y \text{ متقطعة}$

$$(a) f(x,y) \geq 0$$

$$(b) \sum_{y \in R_Y} \sum_{x \in R_X} f(x,y) = 1$$

(2) and if X, Y continuous $\text{عندما } X, Y \text{ متصلة}$

$$(a) f(x,y) \geq 0$$

$$(b) \int \int_{R_X \times R_Y} f(x,y) = 1$$

Continuous Distributions

التوزيعات
المستمرة

Uniform Distribution

① التوزيع المتساوي

$$X \sim U(a, b)$$

$$f(x) = \begin{cases} \frac{1}{b-a} & a \leq x \leq b \\ 0 & \text{otherwise} \end{cases}$$

0

otherwise

الموقع

$$E(X) = \frac{a+b}{2}$$

$$V(X) = \frac{(b-a)^2}{12}$$

التباين

$$M_X(t) = \frac{e^{tb} - e^{ta}}{t(b-a)}$$

, $t \neq 0$, $a \neq b$

مولد العزيم

Normal Distribution

التوزيع الطبيعي

$$X \sim N(M, \sigma^2)$$

$$f(x) = \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{(x-M)^2}{2\sigma^2}} \quad -\infty < x < \infty$$

بذلك ندرس
other wise

$$E(X) = M$$

$$V(X) = \sigma^2$$

الوقع

البات

مولد المزم

$$M_X(t) = e^{Mt + \frac{1}{2}\sigma^2 t^2}$$



Gamma distribution

⑤ توزيع غاما

الشكل العددي
اذا كانت

$$f(x) = \frac{1}{\Gamma(\alpha)} x^{\alpha-1} e^{-x} \quad \alpha > 0$$

فعلنا دالة غاما

$$\Gamma(\alpha) = \int_0^{\infty} x^{\alpha-1} e^{-x} dx = (\alpha-1)!$$

$$\Gamma(\alpha) = (\alpha-1)! \Rightarrow \dots 1$$

$$\Gamma(\alpha+1) = \alpha \Gamma(\alpha) \dots ?$$

$$\Gamma\left(\frac{1}{2}\right) = \sqrt{\pi} \dots ?$$

$$\Gamma\left(\frac{3}{2}\right) = \Gamma\left(\frac{1}{2}+1\right) = \frac{1}{2} \Gamma\left(\frac{1}{2}\right) = \frac{1}{2} \cdot \sqrt{\pi}$$

الحل الثاني: if $X = \frac{Y}{B}$, $B > 0$

$$f(y) = \int \frac{1}{\Gamma(\alpha) B^\alpha} x^{\alpha-1} e^{-\frac{y}{B}} dy$$

$\int_0^\infty$

otherwise

الحل الثاني: دالة توزيع (α, B) لمتغير عشوائي

$$Y \sim G(\alpha, B)$$

$$E(X) = \alpha B$$

$$V(X) = \alpha B^2$$

$$M_X(t) = (1 - Bt)^{-\alpha}$$

مثال: $X \sim G(10, 6)$ دالة التوزيع والخاصة

المعطى

$$f(x) = \frac{1}{\Gamma(10) 6^{10}} x^{10-1} e^{-\frac{x}{6}}$$

$$E(X) = \alpha B = 10 \times 6 = 60$$

$$V(X) = \alpha B^2 = 10 \times (6)^2 = 360$$

$$M_X(t) = \text{دالة}$$

Chi-Square Distribution

مربع کا مربع

$$\beta = 2$$

$$\alpha = \frac{r}{2}$$

If we have a Gamma distribution with $\beta = 2$

and $\alpha = \frac{r}{2}$ where r is positive integer

$$f(x) = \frac{1}{\Gamma(\frac{r}{2})} 2^{\frac{r}{2}} x^{\frac{r}{2}-1} e^{-\frac{x}{2}} \quad x \geq 0$$

2000

this is a chi-square Distribution with (r) degree of freedom $X \sim \chi^2(r)$

$$E(X) = r$$

$$V(X) = 2r \quad \text{and} \quad M_X(t) = (1-2t)^{-\frac{r}{2}}$$

المكانة

مثال

أوجد الدالة الآتية والتبعي والتابعين والدالة $x \in G(10^4, 6^8)$

المولية N و m

$$f(x) = \frac{1}{\sqrt{10} \cdot 6^{10}} \cdot x^{10-1} \cdot e^{\frac{-x}{10}} = \frac{1}{\sqrt{10} \cdot 6^{10}} \cdot x^9 \cdot e^{\frac{-x}{10}}$$

$$E(x) = \alpha B = 10 \cdot 6 = 60$$

$$U(x) = \alpha B^2 = 10 \cdot (6)^2 = 360$$

$$x(t) = \text{داف}$$

Beta Distribution

$$f(x) = \begin{cases} \frac{\Gamma(\alpha + \beta)}{\Gamma(\alpha) \Gamma(\beta)} x^{\alpha-1} (1-x)^{\beta-1} & 0 < x < 1 \\ 0 & \text{otherwise} \end{cases}$$

$$E(x) = \frac{\alpha}{\alpha + \beta}$$

$$V(x) = \frac{\alpha \beta}{(\alpha + \beta)^2 (\alpha + \beta + 1)}$$

$$M_x(t) = \text{مطلوب}$$

$$X \sim \text{Beta}(\alpha, \beta)$$

مطلوب

$$B(\alpha, \beta) = \int_0^1 x^{\alpha-1} (1-x)^{\beta-1} dx$$

دالة بيتا

the probability density function (p.d.f)

دالة الكثافة الاحتمالية

$$f(x) = \frac{1}{B(\alpha, \beta)} x^{\alpha-1} (1-x)^{\beta-1} \quad 0 < x < 1$$

ولا خلاف عن كما

$$\Gamma n = (n-1)!$$

$$\Gamma 3 = (3-1)! = 2! = 2 \cdot 1 = 2$$

قيمة موجبة

$$\Gamma \infty =$$

$$\Gamma 1 = (1-1)! = 0! = 1$$

$$\Gamma 0 = (0-1)! = (-1)! = \infty$$

كل عدد سالب
في ∞ هو ∞

$$\Gamma \frac{1}{2} = \sqrt{\pi}$$

$$\Gamma \frac{3}{2} = \Gamma \left(\frac{1}{2} + 1 \right) = \frac{1}{2} \Gamma \frac{1}{2} = \frac{1}{2} \sqrt{\pi}$$

$$\Gamma \frac{5}{2} = \frac{3}{2} + 1 = \frac{3}{2} \Gamma \frac{3}{2} = \frac{3}{2} \cdot \frac{1}{2} \sqrt{\pi} = \frac{3}{4} \sqrt{\pi}$$

ولا خلاف عن تمام

$$\Gamma n = (n-1)!$$

$$\Gamma 3 = (3-1)! = 2! = 2 \cdot 1 = 2$$

قيمة موجبة

$$\Gamma_{10} =$$

$$\Gamma 1 = (1-1)! = 0! = 1$$

$$\Gamma 0 = (0-1)! = (-1)! = \infty$$

كل عدد سالب
في ∞ هو ∞

$$\Gamma \frac{1}{2} = \sqrt{\pi}$$

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$$\Gamma \frac{5}{2} = \frac{3}{2} + 1 = \frac{3}{2} \Gamma \frac{3}{2} = \frac{3}{2} \cdot \frac{1}{2} \sqrt{\pi} = \frac{3}{4} \sqrt{\pi}$$

Bernoulli dist

$$X \sim \text{Ber}(p)$$

$$f(x) = \begin{cases} p^x q^{1-x} \\ 0 \end{cases}$$

توزيعات

$$x = 0, 1, \dots$$

$$E(X) = p \quad \text{Var}(X) = pq \quad M_X(t) = q + pe^t$$

توزيعات



توزيع ذو الحدين Binomial dist

$$X \sim \text{bin}(n, p)$$

$$f(x) = \begin{cases} \binom{n}{x} p^x q^{n-x} \\ 0 \end{cases}$$

$$x = 0, 1, 2, \dots, n$$

$$E(X) = np, \quad \text{Var}(X) = npq$$

$$M_X(t) = (q + pe^t)^n$$

توزيع بواسون

$$X \sim \text{Po}(\lambda)$$

Poisson distⁿ

$$P(X) = \sum_0^{\infty} \frac{\lambda^x e^{-\lambda}}{x!} \quad x = 0, 1, \dots, \infty$$

$$E(X) = \text{Var}(X) = \lambda \quad M_X(t) = e^{\lambda(e^t - 1)}$$

Negative Binomial distⁿ (5)

$$X \sim \text{Nb}(r, p)$$

$$f(x) = \binom{x+r-1}{r-1} p^r \cdot q^x$$

$$x \geq 0$$

$$0 < p < 1, r > 0$$

$$0 < p < 1, r > 0$$

$$E(X) = \frac{rq}{p}, \quad \text{Var}(X) = \frac{rq}{p^2}$$

$$M_X(t) = \left[\frac{p}{1-qe^t} \right]^r$$

Derivative of Integrals (Leibniz Rule)

$$(1) \frac{d}{dx} \int_a^x f(u) du = f(x) \quad \int_a^x du = x - a$$

$$(2) \frac{d}{dx} \int_{\phi_1(x)}^{\phi_2(x)} f(u) du = f(\phi_2(x)) \frac{d\phi_2(x)}{dx} - f(\phi_1(x)) \frac{d\phi_1(x)}{dx}$$

$$(3) \frac{d}{dx} \int_{\phi_1(x)}^{\phi_2(x)} f(u) du = f(\phi_2(x)) \frac{d\phi_2(x)}{dx} - f(\phi_1(x)) \frac{d\phi_1(x)}{dx}$$

Theorem:- If the random variable $X \sim N(\mu, \sigma^2)$, $\sigma^2 > 0$ then the random variable $W = \frac{X - \mu}{\sigma} \sim N(0, 1)$

Example:- Find $\int_0^{\frac{y+3}{5}} 3x^2 dx$ using Leibniz

Theorem (2) :- If the Random Variable $X \sim N(\mu, \sigma^2)$,
 then the Random Variable $V = \left(\frac{X - \mu}{\sigma}\right)^2 \sim \chi^2(1)$
 مربعی

Proof:-

$$R_X = \{X: -\infty < X < \infty\}$$

$$\text{Range} \left\{ \begin{array}{l} X \rightarrow -\infty \quad \therefore V \rightarrow \infty \\ X \rightarrow \infty \quad \therefore V \rightarrow \infty \end{array} \right.$$

$$R_V = \{V: 0 \leq V < \infty\}$$

$$G(V) = P(V \leq v) = P\left(\left(\frac{X - \mu}{\sigma}\right)^2 \leq v\right) = P(W^2 \leq v)$$

$$= P\{-\sqrt{v} < W < \sqrt{v}\}$$

مربع برعکس 1

$$\int_{-\sqrt{v}}^{\sqrt{v}} f(x) dx$$

$$\frac{1}{\sigma\sqrt{2\pi}}$$

$$= P(\omega: -\sqrt{v} < a < \sqrt{v})$$

$$= \int_{-\sqrt{v}}^{\sqrt{v}} f(x) dx$$

$$= 2 \int_0^{\sqrt{v}} \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}\omega^2} d\omega$$

$$= 2 \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}(\sqrt{v})^2} \cdot \frac{1}{2} v^{-\frac{1}{2}}$$

$$g(v) = \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}(\sqrt{v})^2} \cdot \frac{1}{2} v^{-\frac{1}{2}}$$

$$g(v) = \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}v} \cdot v^{-\frac{1}{2}}$$

مع برهنة
هذا
هو

دالة الكرنج $\sigma w + \mu$

$$G(w) = \int_{-\infty}^{\sigma w + \mu} \frac{1}{\sigma \sqrt{2\pi}} e^{-\frac{1}{2} \left(\frac{x - \mu}{\sigma} \right)^2} dx$$

تملك الطريقة

Leibniz

probability

مساحة الدالة

بفتح
علاوة

$$g(w) = G'(w) = \frac{1}{\sigma \sqrt{2\pi}} e^{-\frac{1}{2} \left(\frac{\sigma w + \mu - \mu}{\sigma} \right)^2} \cdot \sigma$$

$$g(w) = \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2} w^2}$$

$$w \sim N(0, 1)$$

طبيعية

Theorem (2) :- If the Random Variable $X \sim N(\mu, \sigma^2)$
then the Random Variable $V = \frac{(X-\mu)^2}{\sigma^2} \sim \chi^2(1)$
برجسته

Proof:-

$$R_X = \{X: -\infty < X < \infty\}$$

Range

$$\left\{ \begin{array}{l} X \rightarrow -\infty \quad \therefore V \rightarrow \infty \\ X \rightarrow \infty \quad \therefore V \rightarrow \infty \end{array} \right.$$

$$R_V = \{V: 0 \leq V < \infty\}$$

$$G(V) = P(V \leq v) = P\left(\frac{(X-\mu)^2}{\sigma^2} \leq v\right) = P(W^2 \leq v)$$

$$= P\{-\sqrt{v} < W < \sqrt{v}\}$$

مب برسته 1

$$= P \{ w : -\sqrt{v} < a < \sqrt{v} \}$$

$$= \int_{-\sqrt{v}}^{\sqrt{v}} f(x) dx$$

$$2 \int_0^{\sqrt{v}} \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}w^2} dw$$

$$= 2 \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}(\sqrt{v})^2} \cdot \frac{1}{2} v^{-\frac{1}{2}}$$

$$g(v) = 2 \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}(\sqrt{v})^2} \cdot \frac{1}{2} v^{-\frac{1}{2}}$$

$$g(v) = \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}v} \cdot v^{-\frac{1}{2}} \quad \text{--- (1)}$$

منه برهنة ان
الاحتمال
هو 1/2

$$X^2(r) = f(x) = \frac{1}{\sqrt{\frac{\pi}{2}} \cdot 2^{\frac{1}{2}}} \cdot \frac{x^{-1}}{e^{-\frac{x}{2}}}$$

$$X^2(1) \rightarrow f(x) = \frac{1}{\sqrt{\frac{\pi}{2}} \cdot 2^{\frac{1}{2}}} \cdot \frac{x^{\frac{1}{2}-1}}{e^{-\frac{x}{2}}} = \frac{1}{\sqrt{2\pi}} \cdot \frac{x^{-\frac{1}{2}}}{e^{-\frac{x}{2}}} \quad \text{--- ②}$$

مقارنة ① مع ② نجد

$$\nabla \sim X^2(1)$$

$$5) Mx(t) = (1 - \beta t)^{-10}$$

$$X \sim \text{Ga}(\alpha, \beta) \Rightarrow X \sim \text{Ga}(10, \beta)$$

$$Mx(t) = (1 - \beta t)^{-\alpha}$$

$$f(x) = \frac{1}{\Gamma(\alpha) \beta^\alpha} x^{\alpha-1} e^{-\frac{x}{\beta}}$$

$$f(x) = \frac{1}{\Gamma(10) \beta^{10}} x^9 e^{-\frac{x}{\beta}}$$

$$E(X) = \alpha \beta = 10(\beta) = 60$$

$$V(X) = \alpha \beta^2 = 10(\beta^2) = 10(36) = 360$$

Ex: Let X r.v. with p.d.f. $f(x) = 3x$, $0 < x < \sqrt{\frac{2}{3}}$

Let $Y = 3X^2 + 10$ find $g(y)$? give the name of distribution

الكل

$$R_x = \{x : 0 < x < \sqrt{\frac{2}{3}}\}$$

Range

$$\begin{cases} X=0 \rightarrow Y=10 \\ X=\sqrt{\frac{2}{3}} \rightarrow Y=12 \end{cases}$$

$$R_y = \{y : 10 \leq y \leq 12\}$$

prop. $G(y) = P(Y \leq y) = P(3X^2 + 10 \leq y) = P(X \leq \sqrt{\frac{y-10}{3}})$

$$= \int_0^{\sqrt{\frac{y-10}{3}}} 3x \, dx = \frac{3x^2}{2} \Big|_0^{\sqrt{\frac{y-10}{3}}}$$

$$\frac{3}{2} \frac{(y-10)}{3} = \frac{y-10}{2}$$

$$g(y) = G(y) = \frac{1}{2} \Rightarrow \begin{cases} \frac{1}{2} & 10 \leq y \leq 12 \\ 0 & \text{others} \end{cases}$$

$$Y \sim U(10, 12)$$

مقدمات

② one Variable (univariate)

Definition:- let $f(x)$ is ^{دالة} p.m.f. of r.v.s.
and $y = u(x)$ is one to one transformation
then p.m.f of y is ^{مقابل} given by $g(y) = f[\phi(y)]$
where $\phi(y) = x$ is the ^{عكس} inverse of transformation

Ex:- let X is r.v.s. with p.m.f $f(x) = \left(\frac{2}{5}\right), x=1,2,3$
let $y = 3x+2$ find $g(y)$?

$$R_x = \{x : x = 1, 2, 3\}$$

$$R_y = \{y : y = 5, 8, 11\}$$

invers

$$y = 3x + 2 \Rightarrow x = \frac{y-2}{3}$$

عكس الدالة

$$\phi(y) = \frac{y-2}{3}$$

EX3:- let $X \sim \text{bin}(3, \frac{1}{3})$, $Y = X^2$ find $g(y)$?
 $\text{bin}(n, p) \rightarrow P(X) = \binom{n}{x} p^x q^{n-x}$, $x = 0, 1, \dots, n$

$R_X = \{x : x = 0, 1, \dots, n\}$
 $R_Y = \{y : y = 0, 1, 4, \dots, n^2\}$

$Y = X^2 \rightarrow X = \pm \sqrt{y} = \phi(y)$ ناقصة فقط

$g(y) = P(\phi(y)) = P(\sqrt{y})$

$g(y) = \sum_{x \in \phi(y)} \binom{3}{x} \left(\frac{1}{3}\right)^x \cdot \left(\frac{2}{3}\right)^{3-x}$ $q = n - p$
 $y = 0, 1, 4, \dots, n^2$
 other, 1